

Prediction of quantiles by statistical learning and application to GDP forecasting

Joint work with Pierre Alquier, CREST and University Paris 7

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Outline

- 1 Introduction
- 2 The PAC-Bayesian Approach and Gibbs Estimators
 - Gibbs estimators
 - PAC-Bayesian Bound
 - Idea for the proof of PAC-Bayesian Oracle Inequality
- 3 Application to French GDP and quantile prediction

A problem of statistical inference

We observe X_1, \dots, X_n from a \mathbb{R}^p -valued stationary time series $X = (X_t)_{t \in \mathbb{Z}}$ defined on $(\Omega, \mathcal{A}, \mathbb{P})$.

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We have a family of predictors $\{f_\theta : (\mathbb{R}^p)^k \rightarrow \mathbb{R}^p, \theta \in \Theta\}$

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Objective: find a $\theta \in \Theta$ such that $f_\theta(X_{t-1}, \dots, X_{t-k})$ is a good prediction of X_t .

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Definition

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Non-parametric auto-regression predictors:

$$f_\theta(X_{t-1}, \dots, X_{t-k}) = \sum_{i=1}^j \theta_i \varphi_i(X_{t-1}, \dots, X_{t-k}).$$

Loss and risk

Let ℓ be a loss function: $\ell(\hat{X}_t^\theta, X_t) \geq 0$.

Example

- $\ell(x, x') = \|x - x'\|$.
- $\ell(x, x') = \|x - x'\|^2$.

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The (prevision) risk :

$$R(\theta) = \mathbb{E} \left[\ell \left(\hat{X}_t^\theta, X_t \right) \right] \text{ is } \textit{unknown}.$$

The empirical risk:

$$r_n(\theta) = \frac{1}{n-k} \sum_{i=k+1}^n \ell \left(\hat{X}_i^\theta, X_i \right).$$

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Definition (Gibbs Estimators)

We put, for any $\lambda > 0$,

$$\hat{\theta}_\lambda = \int_{\Theta} \theta \hat{\rho}_\lambda(d\theta)$$

where

$$\hat{\rho}_\lambda(d\theta) = \frac{e^{-\lambda r_n(\theta)} \pi(d\theta)}{\int e^{-\lambda r_n(\theta')} \pi(d\theta')}.$$

PAC-Bayesian Bound

The idea is that the risk of the Gibbs estimator will be close to $\inf_{\theta} R(\theta)$ up to a small remainder. More precisely, we upper-bound it by

$$\inf_{\rho} \left\{ \int R(\theta) \rho(d\theta) + \text{remainder}(\rho, \pi) \right\}$$

where the inf is taken upon all the probability distributions on Θ .

Hypothesis

We assume:

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$$\|f_\theta(x_1, \dots, x_k) - f_\theta(y_1, \dots, y_k)\| \leq \sum_{j=1}^k a_j(\theta) \|x_j - y_j\|.$$

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- $k \leq n/2$.

PAC-Bayesian Oracle Inequality

Theorem (Alquier Li 2012)

for any $\lambda > 0$, for any $\varepsilon > 0$,

$$\mathbb{P} \left\{ \int R(\hat{\theta}_\lambda) \leq \inf_{\rho} \left[\int R d\rho + \frac{2\lambda\kappa_n^2}{n(1 - \frac{k}{n})^2} + \frac{2\mathcal{K}(\rho, \pi) + 2 \log(\frac{2}{\varepsilon})}{\lambda} \right] \right\} \geq 1 - \varepsilon$$

where $\kappa_n := \sqrt{2}K(1 + L)(\mathcal{B} + \theta_{\infty, n}(1))$.

θ weak dependent coefficient

Introduced by Doukhan, θ coefficient is as follow:

$$\theta_{\infty}(\mathfrak{G}, Z) = \sup_{h \in \Lambda_1^q} \left\| \mathbb{E}[f(Z)|\mathfrak{G}] - \mathbb{E}[f(Z)] \right\|_{\infty}$$

where

$$\Lambda_1^q = \left\{ f : (\mathbb{R}^p)^q \rightarrow \mathbb{R}, \frac{|f(z_1, \dots, z_q) - f(z'_1, \dots, z'_q)|}{\sum_{j=1}^q \|z_j - z'_j\|} \leq 1 \right\},$$

and that

$$\theta_{\infty, k}(1) := \sup_{p < j_1 < \dots < j_\ell, 1 \leq \ell \leq k} \{ \theta_{\infty}(\sigma(X_t, t \leq p), (X_{j_1}, \dots, X_{j_\ell})) \}.$$

Some preliminary lemmas

Lemma (Rio,2000)

Let h be a function $(\mathbb{R}^p)^n \rightarrow \mathbb{R}$ such that h is 1-Lipschitz. Then for any $t > 0$ we have

$$\mathbb{E} \left(e^{t\{\mathbb{E}[h(X_1, \dots, X_n)] - h(X_1, \dots, X_n)\}} \right) \leq e^{\frac{t^2 n (\mathcal{B} + \theta_{\infty, n(1)})^2}{2}}.$$

Note that $h(X_1, \dots, X_n) = \frac{n-k}{K(1+L)} r_n(\theta)$ and we choose $t = K(1+L)\lambda/(n-k)$, we obtain:

$$\mathbb{E} \left(e^{\lambda[R(\theta) - r_n(\theta)]} \right) \leq e^{\frac{\lambda^2 \kappa^2}{n \left(1 - \frac{k}{n}\right)^2}}.$$

Some proofs

$$\mathbb{E} \left(e^{\lambda \left[R(\theta) - r_n(\theta) - \frac{\lambda \kappa^2}{n \left(1 - \frac{k}{n}\right)^2} \right]} \right) \leq 1,$$

$$\mathbb{E} \left(\int e^{\lambda [R(\theta) - r_n(\theta)] - \frac{\lambda^2 \kappa^2}{n \left(1 - \frac{k}{n}\right)^2} - \log\left(\frac{2}{\varepsilon}\right)} \right) \leq \frac{\varepsilon}{2}.$$

$$\mathbb{E} \left(e^{\sup_{\rho} \left\{ \lambda \int [R(\theta) - r_n(\theta)] \rho(d\theta) - \frac{\lambda^2 \kappa^2}{n \left(1 - \frac{k}{n}\right)^2} - \log\left(\frac{2}{\varepsilon}\right) - \mathcal{K}(\rho, \pi) \right\}} \right) \leq \frac{\varepsilon}{2}.$$

$$\mathbb{P} \left\{ \sup_{\rho} \left\{ \lambda \int [R(\theta) - r_n(\theta)] \rho(d\theta) - \frac{\lambda^2 \kappa^2}{n \left(1 - \frac{k}{n}\right)^2} - \log\left(\frac{2}{\varepsilon}\right) - \mathcal{K}(\rho, \pi) \right\} \geq 0 \right\} \leq \frac{\varepsilon}{2}.$$

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The context

Objective: at each quarter t , predict the flash estimate of GDP growth: ΔGDP_t .

Available information:

- $\Delta\text{GDP}_{t'}$, for all $t' < t$
- $l_{t'}$, for all $t' < t$, l_{t-1} is the climate indicator available to the INSEE at time t .
- The observation period is 1988-Q1 (1st quarter of 1988) to 2011-Q3.

Quantile loss function

We define $X_t = (\Delta\text{GDP}_t, I_t)' \in \mathbb{R}^2$. Following Cornec(CIRET conference 2010), we consider predictors of the form:

$$f_\theta(X_{t-1}, X_{t-2}) = \theta_0 + \theta_1 \Delta\text{GDP}_{t-1} + \theta_2 I_{t-1} + \theta_3 (I_{t-1} - I_{t-2}) |I_{t-1} - I_{t-2}|$$

These family of predictors allow to obtain a forecasting as precise as the INSEE one.

We use the quantile loss function :

$$\begin{aligned} \ell_\tau((\Delta\text{GDP}_t, I_t), (\Delta'\text{GDP}_t, I'_t)) \\ = \begin{cases} \tau (\Delta\text{GDP}_t - \Delta'\text{GDP}_t), & \text{if } \Delta\text{GDP}_t - \Delta'\text{GDP}_t > 0 \\ -(1 - \tau) (\Delta\text{GDP}_t - \Delta'\text{GDP}_t), & \text{otherwise.} \end{cases} \end{aligned}$$

Results: GDP forecasting

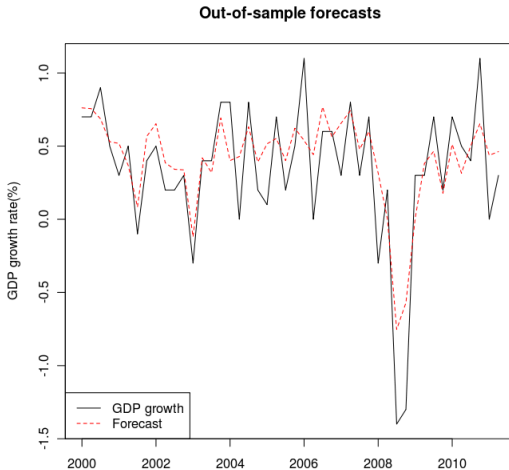


Figure: French GDP online prediction using the quantile loss function with $\tau = 0.5$.

Results: Confidence intervals

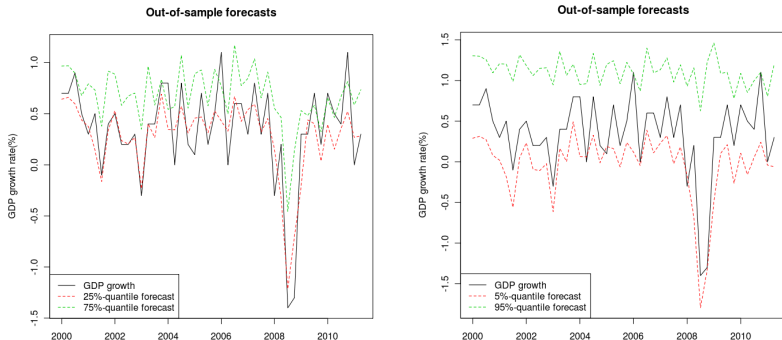


Figure: French GDP online 50%-confidence intervals (left) and 90%-confidence intervals (right).

Conclusion

- We proposed some theoretical results to extend learning theory to the context of weakly dependent time series. The method showed good results on an application to GDP forecasting.
- It would now be interesting to extend this result to a more general class of processes, some other weakly dependent ones.

Merci!